## Age-Related Changes in Objectively Measured Sleep-Wake Are Not Associated With Diurnal Preference: A Big Data Analysis of 18,100 Users

## Introduction

- The circadian system and sleep homeostasis are altered in aging populations¹, yet it remains unclear whether changes in sleep-wake functioning across the lifespan are associated with diurnal preference (i.e., morningness-eveningness).
- Here, we examined whether the slopes of objectively measured sleep changes across the lifespan differed between diurnal preferences.


## Materials \& Methods

Data

- Data from 18,100 users (mean age: 51.4 + 16.6, 58\% female) across 741,738 nights (mean nights recorded: 45) were included in the analysis from the PSGvalidated SleepScore Mobile Application.
- Diurnal preference was subjectively assessed with a 5item questionnaire ranging from definitely morningtype to definitely evening-type.


## Analysis

- Linear mixed effect models were employed to test whether, across age, morningness-eveningness was associated with total sleep time, wake after sleep onset, sleep onset latency, and sleep efficiency.


## Conclusion

- While sleep-wake variables declined linearly with age, the slope of this decline did not differ between strong morning versus strong evening types.
- Age-related sleep impairments are unlikely to be driven by inter-individual differences in morningnesseveningness, despite previous work indicating that diurnal preference reflects dimensions related to circadian periods, sleep homeostasis, and ontogeny.


## Results

|  | Strong Morning | Slight Morning | Neither | Slight Evening | Strong Evening |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Users | $2811(17 \%)$ | $3504(21 \%)$ | $3300(19 \%)$ | $3609(21 \%)$ | $3757(22 \%)$ |
| Age | $51 \pm 16$ | $46 \pm 16$ | $43 \pm 17$ | $41 \pm 16$ | $37 \pm 17$ |
| Female $\%$ | 53 | 60 | 59 | 59 | 60 |
| Nights recorded | 149,965 | 167,038 | 159,301 | 153,487 | 135,342 |
| Bed Time | $22: 37 \pm 1: 14$ | $22: 55 \pm 1: 09$ | $23: 19 \pm 1: 13$ | $23: 43 \pm 1: 13$ | $00: 29 \pm 1: 33$ |
| Wake Up Time | $06: 09 \pm 1: 13$ | $06: 35 \pm 1: 09$ | $06: 59 \pm 1: 15$ | $07: 18 \pm 1: 18$ | $07: 49 \pm 1: 32$ |
| Total Sleep Time (min) | $348 \pm 56$ | $362 \pm 54$ | $363 \pm 58$ | $362 \pm 56$ | $350 \pm 60$ |
| Sleep Efficiency $(\%)$ | $78 \pm 8$ | $80 \pm 8$ | $80 \pm 8$ | $81 \pm 8$ | $81 \pm 8$ |
| Sleep Onset Latency $(\mathbf{m i n})$ | $19.8 \pm 9.5$ | $20.1 \pm 9.7$ | $20.5 \pm 10.0$ | $20.6 \pm 10.3$ | $20.2 \pm 10.3$ |
| Wake After Sleep Onset ( $\mathbf{( m i n})$ | $69.4 \pm 32.9$ | $64.4 \pm 31.6$ | $61.7 \pm 32.0$ | $58.6 \pm 31.3$ | $53.5 \pm 30.0$ |

Table 1. Demographic and average sleep-wake characteristics for all diurnal preferences.


Figures 3 A- 3 D. Although total sleep time (TST, Figure 2A) and sleep efficiency (SEF, Figure 2B) declined linearly with age (TST: $\beta=-1.03, \mathrm{SE}=0.062, \mathrm{p}<0.0001 ;$ SEF: $\beta=-0.27, \mathrm{SE}=0.008, \mathrm{p}=<0.0001$ ), the slopes of this decline were no significantly different between strong-morning and strong-evening types (TST: $\beta=-0.078, S E=0.079, p=0.32$; $\mathrm{SEF}: \beta=-$ $0.001, \mathrm{SE}=0.010, \mathrm{p}=0.91$ ). Similar null findings between strong-morning versus strong-evening types were observed fo sleep onset latency (SOL, Figure 2C) and wake after sleep onset (WASO, Figure 2D), with both increasing linearly with age across all diurnal preferences (WASO: $B=0.40, S E=0.012, \mathrm{p}=<0.0001$; $S O L: B=0.026$, $S E=0.011, \mathrm{p}=0.013$ ), but no significant differences in the slopes over age between strong-morning versus strong-evening types could be observed (WASO: $\beta=-0.005, \mathrm{SE}=0.015, \mathrm{p}=0.74 ; \mathrm{SOL}: ~ \beta=-0.014, \mathrm{SE}=0.014, \mathrm{p}=0.28$ ).

# Assessing the Impact of Race and Income on Changes in Self-Reported Sleep Quality During the COVID-19 Pandemic 

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## Introduction

- Early evidence suggests that the COVID-19 pandemic has differentially impacted sleep-wake functioning¹.
- Here, we examined the association between socioeconomic factors and changes in self-reported sleep quality from before to during the COVID-19 pandemic.


## Materials \& Methods

## Data

- A cross-sectional survey of 2,154 users from the SleepScore database (mean age: 46.8 +- 16.1, 56\% female; 28\% minority race/ethnicity) was conducted in January 2022.


## Analysis

- Proportional odds (ordinal) logistic regression was employed to test the significance of race/ethnicity and annual household income for the likelihood of changes to pre-pandemic self-reported measures of sleep quality, wake after sleep onset (WASO), and sleep onset latency (SOL).


## Conclusion

- Significant changes in self-reported sleep quality during the COVID-19 pandemic were seen across social and economic groups.
- Results suggest that the COVID-19 pandemic may exacerbate pre-pandemic sleep inequalities among individuals with low household incomes.
- A differential impact of the COVID-19 pandemic on selfreported SOL among Hispanic/Latino individuals was observed, though no significant changes to selfreported measures of sleep quality were observed for other racial/ethnic groups.

Results


Figure 1. Hispanic or Latino participants were 1.54 times ( $95 \%$ CI $1.11-2.14, \mathrm{p}=.007$ ) more likely than White participants to report
increased sleep onset latency. No significant changes in self-reported overall sleep quality, soL, or WASO were observed for other increased sleep onset latency. No significant changes in self-reported overall sleep quality, SOL, or WASO were observed for othe racial/ethnic groups.


[^0] $\$ 34,999$ USD were also 1.75 times ( $95 \%$ Cl $1.21-2.53, p=.002$ ) more likely than high-income participants to report decreased \$34,999 USD were


# The Association Between Self-Reported Electronic Device Usage and Objectively Measured Sleep in Adults 

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## Introduction

- Use of light emitting electronic devices in bed before sleep has been associated with sleep disruption in children and adolescents ${ }^{1}$.
- Wavelength and intensity of emitted light, and cognitive and emotional engagement with device, have been proposed as explanations for sleep disruption.
- Here, we examined association between self reported electronic device usage and objective sleep parameters in an adult population of consumer sleep technology users.


## Materials \& Methods

## Data

- Data from 231 users across 25,315 nights from the PSG-validated SleepScore mobile app.
- Device usage was subjectively assessed with a 4item questionnaire ranging from 0 days to 7 days in a week.


## Analysis

- A mixed effect model was used for this analysis.


## Conclusion

- Self-reported use of electronic devices in bed before sleep was associated with shorter time in bed, later bed time and shorter total sleep time.
- These results suggest that electronic device usage before bed reduces the sleep opportunity window and shortens time in bed and total sleep time.


## Results

|  | Full Sample | O Days | $\mathbf{1 - 3 ~ D a y s ~}$ | 4-6 Days | 7 Days |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Number of Users | 231 | 34 | 29 | 70 | 98 |
| Nights Recorded | 25,315 | $3,868(15.2 \%)$ | $5,601(22.1 \%)$ | $5,828(23 \%)$ | $100018(39.6 \%)$ |
| Age (years) | $48.8 \pm 16.6$ | $60.2 \pm 11.1$ | $57.1 \pm 13.9$ | $55.5 \pm 13.7$ | $56 \pm 15.2$ |
| Nights Recorded | - | $15.2 \%$ | $22.1 \%$ | $23 \%$ | $39.5 \%$ |
| Bedtime | $23: 25 \mathrm{hr} \pm 96$ mins | $23: 10 \mathrm{hr} \pm 63$ mins | $23: 45 \mathrm{hr} \pm 96$ mins | $23: 46 \mathrm{hr} \pm 132$ mins | $23: 18 \mathrm{hr} \pm 89 \mathrm{mins}$ |
| Wake Up Time | $7: 25 \mathrm{hr} \pm 120$ mins | $7: 10 \mathrm{hr} \pm 64$ mins | $7: 10 \mathrm{hr} \pm 100 \mathrm{mins}$ | $8: 04 \mathrm{hr} \pm 186$ mins | $7: 17 \mathrm{hr} \pm 88$ mins |
| Total Sleep Time | 352 mins $\pm 75$ mins | 364 mins $\pm 62$ mins | 356 mins $\pm 71$ mins | 341 mins $\pm 80$ mins | 352 mins $\pm 78$ mins |
| Sleep Efficiency (\%) | $76 \pm 12$ | $76 \pm 9$ | $77 \pm 11$ | $76 \pm 11$ | $74 \pm 13$ |



Figure 1. Regression plot of total sleep time for frequency of device use before sleeping (days per week). Higher electronic device
usage was associated with a reduction in total sleep time ( $\beta=-9.2,95 \% C \mid[-15.9-2.5] \mathrm{p}=0.007$.


Figure 2. Regression plots of bed time for frequency of device use before sleeping (days per week). Higher electronic device usage was associated with later bedtimes ( $\beta=0.17,95 \% \mathrm{Cl}[0.029,0.324], p=0.019$ ).



# Alcohol And Caffeine Associated With Poorer Sleep: A Big Data Analysis Of Self-reported Consumption And Objectively Measured Sleep 

## Introduction

- The direct effects of caffeine and alcohol consumption on subsequent sleep have largely been confined to in-lab protocols with cross-sectional measurements and relatively small samples, thus limiting the ecological validity and generalisability of findings.
- This present analysis leveraged longitudinal and naturalistic data from active consumer sleep technology users to examine whether daily self-reported alcohol and caffeine consumption was associated with objectively measured sleep.


## Materials \& Methods

Data

- Data from 26,2448 users across 316,555 nights (mean nights per user: 12.0 +/- 38.8)
- Users aged 16-90 (mean: $38.6+/-15.4$ ) were included in the study. $51.0 \%$ of users were female
- Self-reported questionnaires were used to capture
- Alcoholic beverages consumed
- Caffeine drinks consumed


## Analysis

- Mixed effect modelling was used for the analysis
- Models were adjusted for age and gender


## Conclusion

- Alcohol and caffeine consumption is associated with shorter sleep durations and impaired sleep efficiency, suggesting an overall reduction in sleep quality.
- Alcohol consumption is also associated with a reduction in sleep onset latency
- Our findings suggest that a reduction in alcohol and caffeine consumption by the general public may positively impact sleep health and subsequent general health.


## Results

| Parameter | TST |  |  | SOL |  |  | WASO |  |  | Sleep Efficiency |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef | SE | $\boldsymbol{p}$ | Coef | SE | $\boldsymbol{p}$ | Coef | SE | $\boldsymbol{p}$ | Coef | SE | $\boldsymbol{p}$ |
| Intercept | 419.45 | 1.45 | $<0.001$ | 19.30 | 0.29 | $<0.001$ | 12.52 | 0.64 | $<0.001$ | 91.87 | 0.17 | $<0.001$ |
| Gender[M] | -17.24 | 2.06 | $<0.001$ | 1.49 | 0.41 | $<0.001$ | 3.75 | 0.91 | $<0.001$ | -1.17 | 0.25 | $<0.001$ |
| Age | -0.95 | 0.04 | 0.007 | 0.04 | 0.01 | $<0.001$ | 1.02 | 0.02 | $<0.001$ | -0.24 | 0.00 | $<0.001$ |
| Age:Gender | -0.01 | 0.05 | $<0.001$ | -0.04 | 0.01 | $<0.001$ | 0.07 | 0.02 | 0.001 | -0.02 | 0.01 | 0.002 |
| Drinks | -5.94 | 0.47 | $<0.001$ | -1.02 | 0.12 | $<0.001$ | -0.55 | 0.22 | 0.012 | -0.03 | 0.06 | 0.649 |
| Gender[M]:Drinks | 2.61 | 0.58 | $<0.001$ | 0.02 | 0.15 | 0.905 | 0.75 | 0.27 | 0.006 | -0.08 | 0.07 | 0.254 |
| Age:Drinks | 0.13 | 0.01 | $<0.001$ | 0.01 | 0.00 | 0.021 | 0.03 | 0.01 | $<0.001$ | 0.00 | 0.00 | 0.619 |
| Age:Drinks:Gender[M]] | -0.06 | 0.01 | $<0.001$ | 0.00 | 0.00 | 0.697 | -0.02 | 0.01 | 0.001 | 0.00 | 0.00 | 0.343 |


| Parameter | TST |  |  | SOL |  |  | WASO |  |  | Sleep Efficiency |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef | SE | $\boldsymbol{p}$ | Coef | SE | $\boldsymbol{p}$ | Coef | SE | $\boldsymbol{p}$ | Coeef | SE | $\boldsymbol{p}$ |
| Intercept | 422.29 | 1.60 | $<0.001$ | 18.68 | 0.33 | $<0.001$ | 12.50 | 0.71 | $<0.001$ | 92.03 | 0.19 | $<0.001$ |
| Gender[M] | -16.94 | 2.27 | $<0.001$ | 1.44 | 0.47 | 0.002 | 3.73 | 1.01 | $<0.001$ | -1.10 | 0.27 | $<0.001$ |
| Age | -0.92 | 0.04 | 00.001 | 0.03 | 0.01 | $<0.001$ | 1.04 | 0.02 | $<0.001$ | -0.24 | 0.01 | $<0.001$ |
| Age:Gender | -0.03 | 0.05 | 0.553 | -0.04 | 0.01 | 0.001 | 0.08 | 0.02 | 0.001 | -0.02 | 0.01 | $<0.001$ |
| Cups | -5.82 | 0.59 | $<0.001$ | 0.19 | 0.14 | 0.193 | -0.47 | 0.27 | 0.084 | -0.15 | 0.07 | 0.03 |
| Gender[M]:Cups | 1.28 | 0.76 | 0.093 | -0.11 | 0.19 | 0.57 | 0.31 | 0.35 | 0.386 | -0.05 | 0.09 | 0.576 |
| Age:Cups | 0.06 | 0.01 | $<0.001$ | 0.00 | 0.00 | 0.953 | 0.00 | 0.01 | 0.622 | 0.00 | 0.00 | 0.35 |
| Age:Cups:Gender[M] | -0.01 | 0.02 | 0.407 | 0.00 | 0.00 | 0.844 | -0.01 | 0.01 | 0.171 | 0.00 | 0.00 | 0.178 |



[^1]
[^0]:    Figure 2. Compared to high-income ( $\$ 150,000+$ USD) participants, low-income ( $\$ 0$ - $\$ 25,000$ USD) participants were 2.15 times
    $(95 \% C$ I $1.58-2.92, \mathrm{p}<.001)$ more likely to report decreased

[^1]:    

